



CDO market implosion and the pricing of subprime mortgage-backed securities ☆

Yongheng Deng^{a,*}, Stuart A. Gabriel^b, Anthony B. Sanders^c

^a Institute of Real Estate Studies, National University of Singapore, 21 Heng Mui Keng Terrace, #04-02, Singapore 119613, Singapore

^b UCLA, Anderson School of Management, 110 Westwood Plaza, Entrepreneurs Hall Suite 412, Los Angeles, CA 90095-1481, USA

^c School of Management, George Mason University, 4400 University Drive, Fairfax, VA 22030, USA

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ABSTRACT

We evaluate the effects of CDO issuance on the pricing of subprime residential mortgage-backed securities. Upon controlling for mortgage option values and other well-established determinants of credit spreads, GMM results indicate that the emergence and rapid capitalization of the subprime-backed CDO market was associated with a significant tightening of subprime MBS–Treasury yield spreads. Results of VAR and other robustness tests serve to corroborate the findings. Dynamic simulation based on the impulse response function estimates indicates substantial subprime MBS spread widening in the wake of the recent implosion in the CDO market. Research findings suggest the importance of supply/demand shocks associated with innovations in derivative securities markets to the pricing of securitized subprime debt.

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1. Introduction

The subprime crisis and the housing bubble and its burst have had profound repercussions on the world economy. In particular, the downside of the housing bubble has produced (coupled with rising unemployment) increased foreclosures and further downward pressures on housing prices. One of the contributing factors to the housing bubble was the expansion of credit to subprime borrowers and the introduction of low/no documentation loan products.

The loans to subprime borrowers and the “Alt-A” loans were typically sold to secondary market participants rather than being kept on bank balance sheets. In order to understand the financial crisis, it is important to understand the collateralized debt obligation (CDO) market which allowed the subprime and Alt-A mortgage markets grow dramatically in size.

The global market for collateralized debt obligations (CDOs) witnessed explosive growth over the 1997–2006 period, as the stock of global issuance expanded from \$300 billion to almost \$2 trillion. CDO issuance importantly supported the market for subprime mortgage-backed securities, via the re-packaging of relatively illiquid subordinated assets into derivative CDO securities. In a stunning market about-face, new issuance of CDOs ceased in early 2007, in the wake of implosion and wholesale re-pricing of credit risk in the capital markets. The surge in issuance of subprime-backed CDOs coincided with a marked tightening in subprime MBS–Treasury spreads, suggesting some measurable effect of derivative securitization on the

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* Corresponding author. Fax: +65 6774 1003.

E-mail addresses: ydeng@nus.edu.sg (Y. Deng), stuart.gabriel@anderson.ucla.edu (S.A. Gabriel), asander7@gmu.edu (A.B. Sanders).

Table 1
Residential mortgage deals in 420 ABS CDOs.

Vintage	Subprime	Alt-A	Seconds	Prime	Total
<i>Number of deals by vintage and mortgage loan type</i>					
2003	215	63	7	144	429
2004	371	252	25	188	836
2005	488	452	62	209	1211
2006	522	487	69	142	1220
2007	150	113	21	28	312
Total	1746	1367	184	711	4008

Source: UBS, "Mortgage and ABS CDO Losses," December 13, 2007.

supply/demand balance and pricing of mortgage-backed securities. In 2007 and in the wake of the implosion in the CDO market, spreads on mortgage-backed securities widened markedly.

CDOs are financial structures whereby a set of assets are held in a trust formed as a Special Purpose Vehicle (SPV). A series of tranches with different exposures to the risks of the underlying assets are issued by the trust. A CDO cash-flow structure, for example, allocates the interest and principal payments of the underlying collateral pool of debt instruments to the CDO tranches. While there are many variations, a cash CDO, for example, is a senior-subordinated structure where the senior CDO debt tranches are paid first, then the mezzanine and lower-subordinated notes. Any remaining cash flow is available to equity. In the CDO structure, a set of assets (such as corporate bonds, CMBS, or residential mortgage-backed securities) can be packaged into claims and sold to investors.¹

As shown in Table 1, subprime mortgage asset-backed securities (ABS) overwhelmingly dominated other mortgage products (such as prime and second mortgages) as the collateral for CDOs. As shown in Fig. 1, issuance of subprime asset-backed CDOs moved up fivefold during the first half of the decade—from about \$10 billion in 2000 to in excess of \$50 billion in 2006. Coincidentally, spreads to Treasury on subprime MBS-backed CDOs narrowed substantially. Indeed, as evidenced in Fig. 1, yield spreads of subprime MBS trended down from a high in excess of 500 basis points in 2000 to about two-fifths that level 4 years later. Interestingly, as shown in Fig. 2, that marked trending down in subprime MBS to Treasury spreads occurred even as subprime MBS issuance was rapidly expanding. Increased demand for subprime MBS product, for purposes of derivative securitization via the CDO vehicle, may have resulted, all things equal, in some tightening

of subprime MBS–Treasury spreads. By 2006, however, the underlying collateral for many CDOs, subprime MBS, began to experience sharp increases in delinquencies and defaults. In the wake of sharp erosion in collateral performance, spreads began to widen markedly in both subprime MBS and CDOs and the issuance of CDOs began to decline. By 2007, issuance of CDOs backed by subprime MBS had ceased.

The rise and fall of the CDO market provides an excellent laboratory to study spreads on credit-sensitive securities. Following previous studies on credit spreads (see, for example, Collin-Dufresne et al., 2001; Collin-Defresne and Goldstein, 2001; Chen et al., 2007), we examine the relationship between the spreads on the underlying subprime collateral and the issuance volume of CDOs. In so doing, we build on the prior literature in evaluating the determinants of yields spreads on mortgage-backed securities.

Much of the research on CDOs has focused on pricing of those derivative securities (see, for example, Duffie and Garleanu, 2001; Hull and White, 2003; Noh, 2004; Longstaff et al., 2005; Ericsson and Renault, 2006; Longstaff and Rajan, 2006; Pan and Singleton, 2008). Brennan et al. (2008) evaluate the effects of tranching of debt contracts on the prices at which debt securities can be marketed. In addition, there has been considerable research on a related topic: correlated default among underlying assets (see, for example, Duffee, 1998; Duffie et al., 2009; Giesecke, 2004; Das et al., 2007; Giesecke and Goldberg, 2005).

In discussion of the credit crisis, Gorton (2009) assesses CDO structure and issuance. In particular, Gorton identifies rating-related arbitrage as a primary motivation for CDO issuance.² Further, the opportunity for negative basis trades may have been important to CDO purchase of subprime mortgage-backed bonds.³ Although recent studies have focused on CDO pricing and issuance, little research has attempted to assess the direct impact of CDO market evolution on spreads in the underlying collateral.

Alternative hypotheses can be put forth regarding the effects of the emergence of the CDO market on the pricing of subprime residential asset-backed securities. On the one hand, to the extent CDOs confer efficiency benefits as market-completing investment vehicles, institutional demand for CDOs and related derived-demand for asset-backed securities may increase. According to this *market-completion hypothesis*, CDOs may have served to enhance efficiency in the asset-backed securities markets, via the pooling and re-tranching of less liquid and lower-rated subprime MBS into derivative securities that were more tradable. Further, a liquidity premium associated with senior CDO tranches may have reduced the cost of raising capital through tranches below the cost of acquiring the

¹ While the rules for distributing the cash flows of the CDO bonds' underlying collateral are relatively straightforward, the valuation of the debt and equity tranches can be complicated. The reason for this complexity is that CDOs often contain a large portfolio of assets and credit exposures with diverse risk profiles (for example, default, loss and recovery patterns) as collateral and there are many possible capital structures from which debt and equity are issued. Unlike those mortgage obligations where prepayment-risk is the dominant focus, in CDO valuation the primary focus is typically on the credit risk. A number of practitioner-oriented papers provide excellent discussions of the CDO market (e.g., Goodman and Fabozzi, 2002; Li et al., 2004; Roy and Shelton, 2004; Tavakoli, 2003).

² Gorton (2009) suggests that by 2005, yield spreads on subordinated subprime MBS tranches were elevated relative of other structured products of similar ratings, providing an opportunity to arbitrage the ratings between the ratings on subprime MBS and on the CDO tranches.

³ Such negative basis trades could occur to the extent yields on CDO tranches exceed required payment for credit risk protection (in the CDS market). Originators of structured products had an incentive to engage in ratings-related arbitrage to the extent spreads on subprime MBS tranches exceeded those on similarly rated CDO tranches.

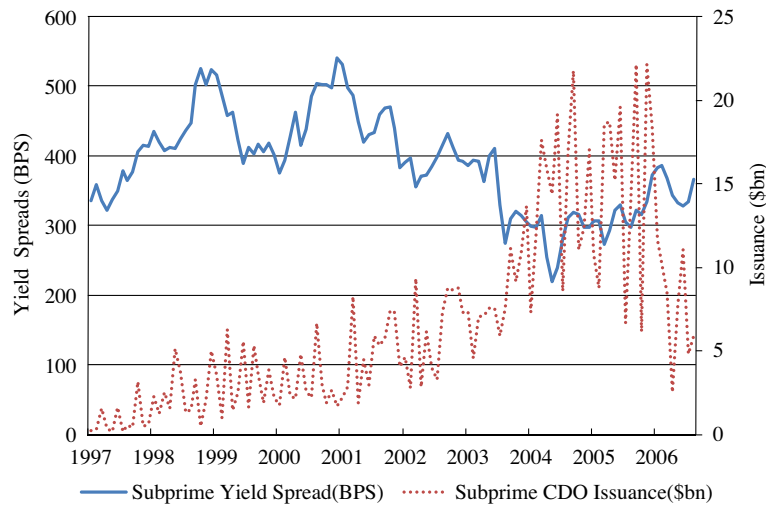


Fig. 1. Yield spread vs. subprime CDO issuance.

asset pool (see, for example, Greenbaum, 1987; DeMarzo and Duffie, 1999; Demarzo, 2003).^{4,5} Assuming less than fully elastic supply of subprime MBS in the capital markets, CDO-related related increments to subprime MBS demand would result in some contraction in mortgage-backed security spreads, all things equal.

On the other hand, capitalization and growth of the CDO market could serve to elicit a supply response, in the form of increased sales of subprime securities in the real estate asset-backed securities markets. Those sales might derive from opportunities for improved price execution and/or enhanced outlets for liquification of investor residential mortgage-backed securities (RMBS) portfolios. As suggested by Greenbaum and Thakor (1987), the advent of new securities markets allows financial intermediaries to remove assets from their balance sheets and reduce other costs of holding debt. All things equal, this *supply-shift hypothesis* would serve to depress subprime asset-backed security prices and concomitantly result in wider MBS–Treasury spreads.

Indeed, prior research has suggested the importance of sector-specific supply/demand imbalances in the determination of yield spreads. Collin-Dufresne et al. (2001) show

that monthly credit spread changes on corporate bonds are largely driven by supply/demand shocks that are independent of both credit-risk factors and standard proxies for liquidity. Duffie and Singleton (1999), controlling for credit-risk and liquidity factors, find that swap market supply/demand shocks drive unexplained changes in swap rates. Similarly, supply/demand shocks appear important to analyses of Ginnie Mae yield spreads (Boudoukh et al., 1997). Evidence also points to the role of supply/demand imbalances in the determination of other residential mortgage yield spreads (see, for example, Bradley et al., 1995; Lehnert et al., 2008).

Alternatively, according to a simple *shell game hypothesis*, elevated demand for subprime-backed MBS might have been driven by inaccurate CDO underwriting or ratings, mis-information, mis-representations, or lack of full understanding of the risks of the CDO vehicle. Some combination of those factors could serve to explain tighter spreads to Treasury on subprime MBS. An alternative *asymmetric information hypothesis* suggests that CDO issuers may possess private asymmetric information regarding asset returns, and engage in related price discrimination via tranching to maximize profits (Oldfield (2000) or use such pooling and tranching methodologies to diversity risk (DeMarzo, 2005).⁶

Finally, one might posit a *production efficiency hypothesis*, whereby specialization and vertical disintegration of such functions as MBS securitization, ownership and servicing, would serve to enhance production efficiency and in so doing decrease MBS production costs. This hypothesis derives from well-known work by Greenbaum (1987) and

⁴ DeMarzo and Duffie (1999) and DeMarzo (2003) build liquidity-based models of tranching. In those models, the security issuer may possess private information regarding security payoff that may cause illiquidity. However, the senior tranches (low risk tranches) are less sensitive to the (CDO) issuer's private information, and thus may enjoy greater liquidity than the underlying collateral.

⁵ On the other hand, DeMarzo (2005) shows that for an informed intermediary, pure pooling and sales of assets from the pool is inferior to selling assets separately. This is because asset pooling eliminates the intermediary's option regarding how aggressively to market each asset and thus can reduce the payoff. This is called the "information destruction effect". However, there can be an offsetting "risk diversification effect" of pooling and tranching – in that the intermediary can create lower-risk derivative securities from the asset pool, and such securities are less sensitive to the intermediary's private information and accordingly can be more attractively priced to the investor. In the case of CDOs, gains from risk diversification were expected to exceed losses from information destruction, such that on net pooling and tranching facilitated higher gains than individual asset sales.

⁶ Oldfield (2000) argues that tranching may allow security issuers to further enhance returns via price discrimination. Assuming that the demand functions for various derivative products are imperfectly price elastic, Oldfield (2000) explains that the security issuer seeks private information about investor demand via the security design and sales process, and uses that information to segment the market and price discriminate among different sets of customers. In the case of CDOs, the price discrimination could be facilitated via the re-bundling of the pool and the selling the different tranches at different prices.

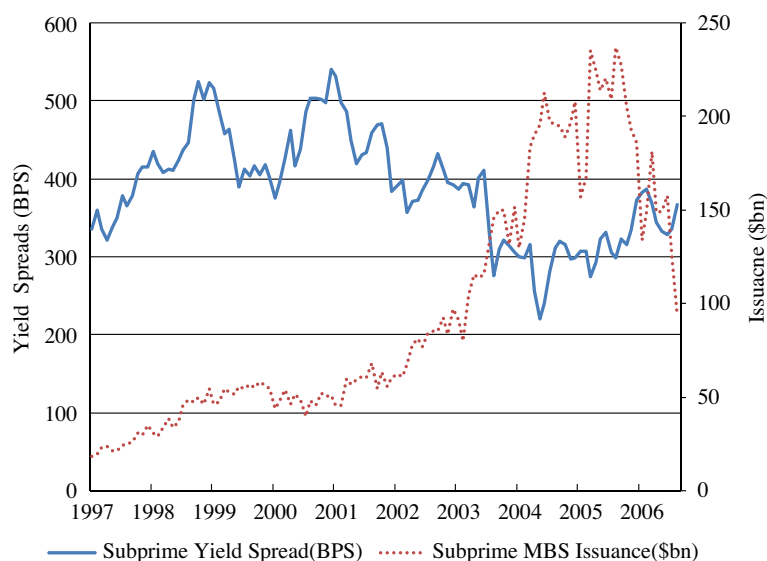


Fig. 2. Yield spread vs. subprime MBS issuance.

Hess and Smith (1988) which posits gains from specialization of activity in each step of the intermediation and securitization function. Here we would similarly anticipate some contraction in MBS–Treasury spreads.

In this paper, we undertake empirical assessment of the effects of CDO issuance on the pricing of subprime MBS. We employ various empirical specifications to identify CDO issuance effects. Firstly, to address causality, we undertake Granger Causality Wald Tests. Results of Granger Tests (see Table 2) support the hypothesis that issuance of subprime-backed CDOs Granger cause changes in subprime–Treasury spreads, but not vice versa.⁷ In order to disentangle the economic relationship between subprime–Treasury spreads and issuance of subprime-backed CDO along with other risk factors, we undertake parametric estimation of subprime MBS pricing determinants using GMM models with Newey–West Kernel error correction. Those models specify and control for well-established options embedded in MBS–Treasury spreads. Finally, to further explore the dynamic dependency between subprime–Treasury yield spreads and the underlying time-series variables, we estimate a structural vector autoregression (VAR) model with changes in subprime–Treasury yield spreads and issuance of subprime-backed CDOs as endogenous variables. The estimated impulse responses from the VAR model are then used to dynamically simulate the effects of implosion of derivative CDO markets on the pricing of subprime mortgage-backed securities.

Results of GMM model estimation suggest that factors associated with the termination risks of the underlying subprime residential mortgage contracts, including interest-rate volatility, the term structure of interest rates,

Table 2
Granger Causality Wald Test.

Null hypothesis	F-statistic	P-value
In CDO issuance does not Granger cause Δ Subprime spread	3.1555	0.0279*
Δ Subprime spread does not Granger cause In CDO issuance	0.1696	0.9167

* Indicates rejection of the null hypothesis at the 5% level.

and proxies for credit risk in the macro-economy, importantly affect the magnitude of mortgage security–Treasury spreads. Further, subprime MBS to Treasury spreads evidence only limited sensitivity to returns and related return volatility among alternative asset classes, notably including those of equity markets. Finally, holding constant those factors, research indicates that the emergence of the CDO market was associated with a significant contraction in subprime mortgage security to Treasury yield spreads.

VAR estimation shows that CDO issuance lagged 2 months exerts significant, negative effects on subprime MBS–Treasury spreads. Also, as a further indication of the absence of reverse causality, none of the lagged coefficients for changes in subprime MBS–Treasury spreads are statistically significant in the estimated impulse response of CDO issuance. We then use the estimated impulse response functions from the VARs to simulate changes in subprime mortgage-backed security pricing associated with cumulative negative shocks in the issuance of subprime-backed CDOs. Results indicate that a one standard deviation negative shock to In CDO issuance over a period of 3, 6, and 9 months, results in a widening of subprime MBS–Treasury yield spreads by 110 bps, 150 bps, and 170 bps, respectively. Further, findings suggest that a one standard deviation positive shock in subprime MBS–Treasury spreads lasting for 3–9 months leads to further widening in spreads by up to 1000 bps. The latter captures

⁷ Table 2 reports the Granger Causality Wald Test statistics. The *p*-value < 5% allows rejection of the null-hypothesis that the natural log of issuance of subprime-backed CDOs does not Granger Cause changes in subprime–Treasury spreads.

the dynamic process of dramatically widening in subprime MBS–Treasury spreads during the recent financial crisis.

Note further that CDO pricing effects likely were passed back to mortgage borrowers in the primary market. In that regard, interest rates faced by subprime borrowers rose markedly in the wake of the 2007 collapse of CDO markets and related reductions in the valuation of subprime mortgage pools. Accordingly, our results provide systematic evidence linking derivative CDO markets to the pricing of underlying subprime mortgages.

The remainder of this paper is organized as follows: Section 2 discusses the influence of CDOs on the pricing of subprime residential mortgage-backed securities. Section 3 describes the data used in the study and presents the empirical model, and Section 4 presents the empirical findings and simulation results. Section 5 summarizes the results and provides concluding remarks.

2. The role of CDOs in the determination of subprime MBS yield spreads

The termination risks of mortgage-backed securities differ substantially from those of US Treasuries. Treasury obligations provide a full faith and credit US government guarantee of timely repayment of principal and interest. In marked contrast, the cash flows and hence pricing of mortgage-backed securities reflect regularities associated with borrower exercise of mortgage put and call options. Those borrower options are typically written into the mortgage contract; in well-functioning markets, option valuations are embedded in the mortgage–Treasury rate spread. In the case of the mortgage put option, mortgage borrowers may put the mortgage back to the lender in the case of default. Also, call option exercise is often permitted in the case of residential mortgages. Indeed, the inadequacy of both underwriting and pricing of default risk over past years has been fundamental to the subprime crisis.⁸

Frequent borrower exercise of the mortgage call option, in the form of mortgage prepayment, was evidenced over the period of analysis.⁹ Indeed, prepayment risks should vary with expectations regarding the future path of interest rates. Also, higher levels of interest rate volatility suggest higher call option values and elevated risks of mortgage termination. Those risks should be priced into higher mortgage spreads over Treasuries. However, this effect is mitigated somewhat in environments with steeper yield curves, as this term structure suggests that interest rate declines associated with volatility will be more muted, thereby implying diminished termination risk premia for seasoned mortgage product.

⁸ In subprime residential debt markets, the MBS issued by investment banks were backed by private mortgage insurance and other forms of credit enhancement rather than the full faith and credit of the US government.

⁹ The recent housing boom years witnessed a secular increase in residential mortgage prepayment speeds, owing to technological innovation and reductions in the costs of mortgage re-finance, enhanced mortgage product offerings, improved borrower knowledge of mortgage financing, ease of loan qualification, and substantial downward adjustment in mortgage interest rates subsequent to the 2001 recession. In the case of subprime mortgages, conditional on adequate house price appreciation, refinancing was common over relatively short time horizons.

Spreads between mortgage and Treasury securities also may reflect variability among asset classes in risk-adjusted returns. As such, changes in investor asset allocation among mortgage and corporate debt, equity, and government bonds could markedly affect the pricing of mortgage-backed securities. For example, higher returns to equities might result in investor portfolio re-allocation to that asset class, resulting in damped demand for and elevated premia among securitized mortgage product. Of course, elevated equity returns likely would have less influence on portfolio asset allocation to the extent the higher equity returns are accompanied by elevated return volatility.

Finally, numerous hypotheses can be brought to bear as regards the effects of rise and fall of the CDO market on the pricing of subprime MBS. As suggested above, the *market-completion, shell game, specialization and production*, and *asymmetric information hypotheses* all point to a tightening of subprime MBS–Treasury spreads in the wake of the emergence and rapid capitalization of the CDO market. In contrast, the *supply-shift hypothesis* suggests that growing capitalization of the CDO vehicle may have elicited a MBS supply response and hence a widening of spreads to Treasuries. While the relative magnitudes of those influences are difficult to entangle, we test below for the influence of CDO issuance in the determination of subprime MBS–Treasury spreads, controlling for the relative magnitude of CDO to subprime MBS issuance and other well-established pricing determinants.

2.1. Theoretical determinants of mortgage bond spreads

2.1.1. Slope of the Treasury yield curve

There exists substantial evidence on the role of the term structure in the determination of mortgage bond spreads (see, for example, Bradley et al., 1995; Ambrose and Sanders, 2003; Titman et al., 2004). As is widely-appreciated, an increase in the slope of the yield curve suggests some future strengthening in economic activity. As such, a steeper yield curve imparts a higher probability of a short rate increase. A steeper yield curve reduces the likelihood that the mortgage call option will be in the money, so as to reduce prepayment risk and the related call option premium. Accordingly, increases in the slope of the Treasury yield curve should have a negative impact on the mortgage-backed security–Treasury spread.

2.1.2. Volatility

Mortgage put and call option values increase with interest rate volatility. In fact, in a contingent claims framework, the debt claim has elements similar to a short position on a put option. Mortgage bond spreads should increase with volatility. This prediction is intuitive and is well established in the literature; increased interest rate volatility implies higher values for embedded options hence increases in the probability of prepayment and default.

2.1.3. Credit risk

Fama and French (1989) find that credit spreads widen when economic conditions are weak. We use the interest rate spread between corporate bonds rated Aaa and Baa to proxy such economy-wide credit risk. In the current

Table 3

Explanatory variables and expected signs of the coefficients of the basic model.

$$\Delta(r_{B_t} - r_{N_t}) = \alpha + \beta_1 \Delta Slope_t + \beta_2 \Delta Vol_{CMT_t} + \beta_3 \Delta Slope_t \times \Delta Vol_{CMT_t} + \beta_4 \Delta(r_{Aaa_t} - r_{Baa_t}) + \beta_5 \Delta S\&P_t + \beta_6 \Delta S\&P_t \times \Delta Vol_{S\&P_t} + \beta_7 \ln(CDO_Iss_t) + \beta_8 \frac{CDO_Iss_t}{Section_Iss_t} + \mu_t$$

Variable	Description	Predicted Sign
$\Delta(r_{B_t} - r_{N_t})$	Change in spread between subprime MBS yield and 10-year CMT yield (dependant variable)	
$\Delta Slope$	Change in spread between 10-year CMT and 3-month CMT yields	–
ΔVol_{CMT}	Change in 10-year CMT volatility	+
$\Delta Slope \times \Delta Vol_{CMT}$	Interaction of change in slope and change in volatility	–
$\Delta(r_{Aaa} - r_{Baa})$	Change in default risk measured by spread in composite yields on corporate Aaa and Baa bonds rated by Moody's	+
$\Delta S\&P$	Change in excess return on S&P 500 (dividend included) index over the 3-month CMT	+
$\Delta S\&P \times Vol_{S\&P}$	Interaction of change in S&P 500 return and S&P 500 volatility	–
$\ln(CDO_Iss)$	Natural log of dollar volume of subprime CDO issuance	–
$\frac{CDO_Iss}{Section_Iss}$	Ratio of subprime CDO issuance over subprime residential MBS issuance	+

application, we hypothesize that the put option embedded in the mortgage contract should vary directly with economy-wide credit risk and the market price of risk. The credit spread proxy has been previously applied to proxy mortgage default risk (see, for example, [Duca and Rosenthal, 1991](#); [Bradley et al., 1995](#)). Further, recent research by [Anginer and Yildizhan \(2008\)](#) shows that corporate credit spreads do a better job of predicting defaults than previously used measures, such as bond ratings or accounting based parameters.

2.1.4. *Alternative assets returns*

As evidenced in [Kwan \(1996\)](#) and [Collin-Dufresne et al. \(2001\)](#), returns to alternative asset classes, notably including equity returns, may affect the demand for and yields on fixed income securities. Similar to those papers, we include excess return in the S&P 500 over the 3-month Constant Maturity Treasury yield as well as a measure of volatility of S&P 500 excess returns to proxy returns on alternative equity investment classes.

2.2. *CDO market evolution: determinants related to market capitalization and supply/demand imbalances*

2.2.1. *Subprime MBS-backed CDO issuance*

Increases in subprime-backed CDO issuance associated with market completion, specialization and production efficiency, asymmetric information, or shell game effects should serve to elevate demand for underlying subprime securitized product. All things equal, such increments to demand for subprime MBS should result in a tightening of subprime MBS–Treasury spreads. By the same token, negative shocks to demand for subprime MBS, due to an implosion in subprime-backed CDO issuance, should have the opposite effect.

2.2.2. *Subprime MBS-backed CDO issuance/subprime MBS issuance*

We further control for liquidity effects using the ratio of subprime-backed CDO to subprime MBS issuance. Changes in that ratio could arise, for example, owing to differences across secondary and derivative markets in recognition of

performance problems associated with underlying subprime mortgage collateral and related adverse impacts on security issuance and liquidity of trades. For example, collateral performance problems could lead to dampened liquidity of trades and diminished issuance of subprime MBS. However, to the extent opaqueness of CDO structures and over-reliance on ratings agencies served to mask and hence delay CDO investor recognition of such performance problems, one would expect the ratio of subprime CDO to subprime MBS issuance to exert upward pressure on spreads.¹⁰

Table 3 provides description of the explanatory variables and summarizes the predicted sign of the changes in subprime–Treasury spreads in response to the changes in the underlying determinants.

3. **GMM model and data**

We estimate the below reduced form model.

$$\Delta(r_{B_t} - r_{N_t}) = \alpha + \beta_1 \Delta Slope_t + \beta_2 \Delta Vol_{CMT_t} + \beta_3 \Delta Slope_t \times \Delta Vol_{CMT_t} + \beta_4 \Delta(r_{Aaa_t} - r_{Baa_t}) + \beta_5 \Delta S\&P_t + \beta_6 \Delta S\&P_t \times \Delta Vol_{S\&P_t} + \beta_7 \ln(CDO_Iss_t) + \beta_8 \frac{CDO_Iss_t}{Section_Iss_t} + \mu_t \tag{1}$$

In the above equation, the dependent variable is measured as the basis point differential in spreads between the relevant mortgage and Treasury securities. All models are estimated in first differences using the GMM. The Newey–West Kernel approach is used for error correction. The independent variables in the estimating equations are as described above. Our specification is consistent with well-developed analyses of the determinants of changes in credit spreads (see, for example, [Duffee, 1998](#); [Collin-](#)

¹⁰ See article titled “CDO Boom Masks Subprime Losses, Abetted by S&P, Moody’s, Fitch” as published in *Bloomberg News, May 31, 2007*. In that article, it was reported that “CDO holdings have already declined in value between \$18 billion and \$25 billion because of falling repayment rates by subprime US mortgage holders.” The article went on to say that “In many cases, investors don’t even know that values have dropped.”

Table 4
Descriptive statistics and unit root test.

	Descriptive statistics				Unit root test (<i>P</i> -value)	
	Mean	Maximum	Minimum	Std. dev.	w/o trend	w/ trend
Subprime spread	3.8868	5.4080	2.2030	0.6986	0.3899	0.2365
Slope	1.4410	3.6800	−0.7000	1.2137	0.6302	0.9326
Vol_{CMT}	0.2182	0.4550	0.0758	0.0814	0.0065*	0.0334*
$Slope \times \Delta Vol_{CMT}$	0.3679	1.3220	−0.1300	0.3764	0.5076	0.8290
Default risk	0.0086	0.0141	0.0055	0.0022	0.1812	0.4197
S&P	0.06363	1.4619	−2.7889	0.8078	0.0000*	0.0000*
$S\&P \times \Delta Vol_{S\&P}$	−0.0024	0.5480	−1.1215	0.2344	0.0000*	0.0000*
$\ln(CDO_Iss)$	1.4573	3.0978	−1.6663	1.0660	0.2279	0.2430
$CDO_Iss/Section_Iss$	4.3237	5.4679	2.9113	0.6893	0.2833	0.9934

Note: The last two columns give the Phillips–Perron test of the null hypotheses that the indicated series have a unit root.

* Indicates statistical significance at the 5% level.

Dufresne et al., 2001). In that regard, our models account for well-established empirical proxies for put and call option values embedded in the mortgage contract (see, for example, Bradley et al., 1995). Further, as is commonplace in the credit spread literature, our specification controls for excess returns among alternative asset classes. Finally, as suggested above, of central interest to this analysis is the effect of issuance of subprime-backed collateralized debt obligations on the pricing of the underlying subprime bonds. We estimate those effects over two sample periods.

3.1. Data

Our study required access to data on subprime-backed CDO issuance as well as similar monthly-frequency information on subprime MBS issuance and pricing. As suggested below, ABAlert, the primary repository of information on subprime-backed CDOs, provided information on issuance of domestic US subprime-backed cash CDOs. TrueStandings securities provided monthly time-series—based on information from the LoanPerformance dataset—on subprime mortgage-backed security issuance and pricing. Finally, a variety of more readily available data sources were utilized to control for equity returns and the like. A detailed discussion of data resources is provided below. Our sample covers the period from January 1997 to August 2006.¹¹ The descriptive statistics for our data is presented in Table 4. In addition to the mean, maximum, minimum and standard deviation for each variable, we also perform a unit root test both with and without a trend.¹²

¹¹ The timeframe of the analysis corresponds to the rise and fall of the subprime-backed CDO market.

¹² The unit root test statistics reported in Table 4 indicate that many of the time series measured in levels have a unit root. As discussed by Granger and Newbold (1974), stochastic trends (which have a unit root) can lead two time series to appear to be related when in actuality they are not. To avoid potential estimation bias due to spurious correlation, we use the first difference of all the time series except for \ln of CDO issuance and the ratio of subprime CDO to subprime MBS issuance in our regressions. Results of further tests indicate that the first difference in the time series variables are significant at the 1 percent level in our unit root tests, with the exception of the mortgage-backed security–Treasury yield spreads, y_t , which is significant at the 10% level. As a robustness check to address the potential estimation bias with a marginally significant yield spreads process, y_t , we also test our model by replacing yield spreads with the ratio of yield spreads to the 10-year Treasury yield, y'_t , such that $y'_t = X_t\beta + \mu_t$, where y'_t is significant at 1% level in our unit root test.

Data sources and variable construction are as follows:

3.1.1. Yield spreads on asset-backed bonds

Subprime residential MBS yields are obtained from TrueStandings securities. The subprime rates are the weighted average of all residential mortgages classified as subprime by LoanPerformance. Subprime mortgage-backed bond yield spreads are then defined as the difference between the stated bond yield and the associated yield of the constant maturity 10-year Treasury.

3.1.2. Slope of the treasury yield curve

The Treasury yield data—yields on the 10-year Constant Maturity Treasury (CMT) and spread between the 10-year CMT and the 3-month CMT—are obtained from the Federal Reserve Economic Data website at the Federal Reserve Bank of St. Louis. These data are published as monthly averages of their respective series. We interpret this variable as both an indicator of expectations of future short rates as well as an indicator of overall status of the economy.

3.1.3. Volatility

Our measure of interest rate volatility is the annualized standard deviation of the log differences in daily yields on the 10-year Treasury. We also compute the interaction of the yield curve and volatility terms.

3.1.4. Credit risk

The monthly time-series on corporate bonds rates Aaa and Baa are obtained from Moodys Investors Services. This variable is defined as the yield on corporate bonds rated Baa minus the yield on corporate bonds rated Aaa.

3.1.5. Equity market returns

The monthly S&P 500 return series including dividends and price changes is obtained from Datastream. We use the S&P 500 return series to measure both equity market returns and related volatility in that return.

3.2. CDO market evolution

3.2.1. Subprime CDO issuance

We obtain information on CDO issuance from ABAlert.com. The ABAlert.com database contains information on

Table 5

GMM estimations of subprime yield to Treasury spreads on subprime CDO issuance.

$$\Delta(r_{B_t} - r_{N_t}) = \alpha + \beta_1 \Delta Slope_t + \beta_2 \Delta Vol_{CMT_t} + \beta_3 \Delta Slope_t \times \Delta Vol_{CMT_t} + \beta_4 \Delta(r_{Aaa_t} - r_{Baa_t}) + \beta_5 \Delta S\&P_t + \beta_6 \Delta S\&P_t \times \Delta Vol_{S\&P_t} + \beta_7 \ln(CDO_Iss_t) + \beta_8 \frac{CDO_Iss_t}{Section_Iss_t} + \mu_t$$

	Model 1 (1997–2006)	Model 2 (2001–2006)
Constant	–0.0494 (–1.51)	–0.0813 (–1.03)
$\Delta Slope$	–0.7567 (–10.03)	–0.8086 (–7.83)
ΔVol_{CMT}	1.1789 (3.76)	0.6377 (0.84)
$\Delta Slope \times \Delta Vol_{CMT}$	–0.4918 (–4.20)	–0.3160 (–1.22)
$\Delta Default Risk$	0.1712 (1.33)	0.1018 (0.69)
$\Delta S\&P$	0.0295 (0.82)	0.0199 (0.45)
$\Delta S\&P \times \Delta Vol_{S\&P}$	–0.1371 (–1.11)	–0.1940 (–1.58)
In of subprime CDO issuance	–0.0457 (–2.88)	–0.1295 (–3.21)
Ratio of subprime CDO issuance to subprime mortgage issuance	1.4117 (2.49)	4.4379 (4.62)
Number of observations	113	67
Adjusted R^2	0.5610	0.5319

Notes: (1) All models are estimated by GMM approach. Newey–West Kernel is used for error corrections. Dependent variables are the first difference of spreads between the yield on subprime mortgage debt and the 10-year Treasury CMT. Model 1 is estimated based on full sample of data from 1997.1 to 2006.8. Model 2 is estimated based on a sub-sample of data from 2001.2 to 2006.8. (2) $\Delta Slope$ is change in spread between the 10-year CMT and the 3-month CMT. ΔVol_{CMT} is change in the volatility of 10-year CMT. $\Delta Default Risk$ is measured by the change in the spread between the yields on long term Aaa and Baa corporate bonds. $\Delta S\&P$ is change in excess return of S&P 500 (dividend included) index over the 3-month CMT. $\Delta Vol_{S\&P}$ is change in the volatility of excess return of S&P 500 index over the 3-month CMT. Subprime CDO issuance is measured in billions of dollars. (3) t -Statistics are in parenthesis.

the initial terms of all rated issuance of asset-backed securities, mortgage-backed issues and collateralized debt obligations.¹³ The database provides the information on CDO issuance by all the major players, including Wells Fargo, ABN AMRO/LaSalle Bank, J.P. Morgan Chase, and the like. The ABAlert.com database assigns each issue to one of five categories: public ABS, private ABS, MBS, non-US ABS and CDOs. To compute our subprime CDO issuance series, we aggregate the dollar volume of domestic, cash CDOs labeled as subprime CDOs by ABAlert. In order to qualify as a subprime CDO, ABAlert required that at least 75% of underlying collateral was comprised of subprime MBS.¹⁴ The full data timeframe for subprime CDO issuance is from January 1997 to August 2006.

3.2.2. Subprime CDO issuance/subprime MBS issuance

Ratio of subprime MBS-backed CDO issuance to subprime MBS issuance, subprime residential MBS issuance is obtained from TrueStandings securities.

¹³ The primary objective of the ABAlert.com database is to identify the primary participants in each transaction. This database does not include information on pricing or other tranche-specific information. The database captures only the terms of each issue as of its pricing date, so it does not reflect subsequent events, such as paydowns and rating changes. (For further information, see: http://abalert.com/NewPages/Index.cfm?Article_ID=41086).

¹⁴ In computing the dollar volume of CDO issuance, we exclude issuance related to bank loans (CLOs).

4. Estimation and simulation results

4.1. Static analysis

Results from the static analysis based on GMM estimation of subprime asset-backed security–Treasury spreads equations are summarized in Table 5.¹⁵ Table 5 contains estimates for Model 1: 1997–2006 and Model 2: 2001–2006. As evidenced in Table 5, the estimated coefficients are of expected sign. For example, in Model 1, our measure of interest rate volatility, the change in the annualized standard deviation of the log differences in daily yields on the 10-year Treasury, is positive and highly significant in the determination of changes in MBS pricing. Overall, results coincide with priors that higher levels of interest rate volatility elevate the likelihood of prepayment and default option exercise. Accordingly, increases in interest rate volatility are positively priced into mortgage–Treasury spreads. As expected, the term reflecting changes in the slope of the Treasury yield curve enters the mortgage–Treasury spread equations with a negative and significant coefficient, as a steeper yield curve reduces the likelihood that the mortgage call option will come into the money, so as to reduce prepayment risk and the related call option premium.

¹⁵ In other analysis, we also tested the sensitivity of results to comparable-maturity swap spreads. Findings were robust to alternative specifications of the underlying yield curve proxy.

Table 6

Coefficient estimates of the impulse response models.

$$\begin{cases} x_t = \alpha_{10} + \beta_{12}y_t + \sum_{i=1}^3 \delta_{1i}x_{t-i} + \sum_{j=1}^3 \varphi_{1j}y_{t-j} + \gamma B + u_{xt} \\ y_t = \alpha_{20} + \sum_{i=1}^3 \delta_{2i}x_{t-i} + \sum_{j=1}^3 \varphi_{2j}y_{t-j} + \gamma B + u_{yt} \end{cases}$$

Variable	ΔSub Spread	ln CDO Iss	ΔSub Spread	ln CDO Iss
	(1997.01–2006.08)		(2001.02–2006.08)	
Constant	0.0137 (0.3596)	-0.6674 (-8.4744)	-0.0997 (-0.8943)	-0.5794 (-2.8604)
ΔSlope (t)	-0.5691 (-8.7505)	-0.2228 (-1.6624)	-0.6285 (-5.7241)	-0.4260 (-2.1349)
ΔVol _{CMT} (t)	1.5386 (4.6273)	-0.1976 (-0.2884)	0.4433 (0.5856)	-0.8440 (-0.6135)
ΔSlope × ΔVol _{CMT} (t)	-0.5912 (-3.7823)	0.1153 (0.3580)	-0.1962 (-0.6850)	0.4265 (0.8194)
ΔDefault Risk (t)	0.1161 (0.6004)	0.2054 (0.5153)	-0.0250 (-0.1035)	0.0022 (0.0049)
ΔS&P (t)	0.0032 (0.0943)	-0.0807 (-1.1678)	-0.0124 (-0.2404)	0.0177 (0.1891)
ΔS&P × ΔVol _{S&P} (t)	-0.0211 (-0.1921)	0.2576 (1.1359)	-0.0066 (-0.0416)	-0.0172 (-0.0600)
Ratio of CDO Iss (t)	0.4741 (0.9699)	15.9770 (15.8619)	1.8023 (2.3201)	14.1596 (10.029)
ln of CDO Iss (t)	-0.1060 (-1.1221)	-	-0.1353 (-1.1076)	-
ΔSub Spread (t - 1)	0.0324 (0.4698)	-0.0618 (-0.4348)	0.0511 (0.5132)	-0.2242 (-1.2384)
ln of CDO Iss (t - 1)	0.0437 (1.9480)	0.1805 (3.9034)	0.0414 (0.9681)	0.2483 (3.1979)
ΔSub Spread (t - 2)	-0.1653 (-2.6635)	-0.1460 (-1.1412)	-0.2642 (-3.1549)	-0.0890 (-0.5851)
ln of CDO Iss (t - 2)	-0.0727 (-3.4741)	0.2693 (6.2476)	-0.0560 (-1.4131)	0.2481 (3.4447)
ΔSub Spread (t - 3)	0.0300 (0.4885)	0.0788 (0.06219)	-0.0033 (-0.0385)	-0.0809 (-0.5186)
ln of CDO Iss (t - 3)	-0.0062 (-0.2647)	0.3059 (6.3309)	-0.0156 (-0.3658)	0.2893 (0.0773)
Adjusted R ²	0.6460	0.9082	0.6433	0.7847
Akaike AIC	-0.8957	0.5504	-0.7741	0.4208

Notes: (1) All models are estimated using VAR(3), where x and y are endogenous variables and B is a set of exogenous variables. Endogenous variables include subprime spread and ln of CDO issuance; exogenous variables include ΔVol_{CMT} ; $\Delta Slope$; $\Delta Slope \times \Delta Vol_{CMT}$; $\Delta S\&P \times \Delta Vol_{S\&P}$; $\Delta S\&P$; $\Delta Default Risk$; ratio of CDO issuance to subprime issuance. (2) t -Statistics are in parenthesis.

Note that the interaction of the slope of yield curve and volatility change terms enters the model with the expected negative coefficient, suggesting significantly diminished volatility-related widening of the MBS–Treasury spread as the yield curve steepens. Finally, in Model 1, while the default risk proxy and the alternative asset return have the expected positive effect on changes in the MBS spread, neither coefficient is precisely estimated.

Also as anticipated our proxy for the ratio of subprime CDO to subprime ABS issuance is positive and significant throughout. As subprime MBS issuance slows down due to adverse performance in the underlying subprime mortgage collateral, continued CDO issuance relative to a declining subprime MBS asset base should lead to an increase in spreads.

Finally, all things equal, research findings suggest that the rapid growth in CDO issuance has been significant to the pricing of subprime residential mortgage-backed securities. Results here are consistent with CDO-related incre-

ments to demand for subprime MBS associated with market completion, shell game, asymmetric information, and production efficiency hypotheses. As evidenced in Table 5, the CDO issuance coefficient is negative and statistically significant; further, as expected, the estimated magnitude of the CDO issuance effect approximately triples in value in the 2001–2006 period of explosive growth in CDO market capitalization.¹⁶ Furthermore, as anticipated, the ratio of subprime CDO issuance to subprime MBS issuance enters the model with a negative and significant coefficient. Also, the estimated coefficient was substantially larger

¹⁶ We also test the model using data on CMBS tranches spreads to swaps. All things equal, research findings suggest that the rapid growth in commercial real estate-backed CDO issuance has been significant to the pricing of CMBS. The CDO issuance coefficient is negative and statistically significant throughout; further, as expected, the estimated magnitude of the CDO issuance coefficient increases monotonically with reductions in tranche rating.

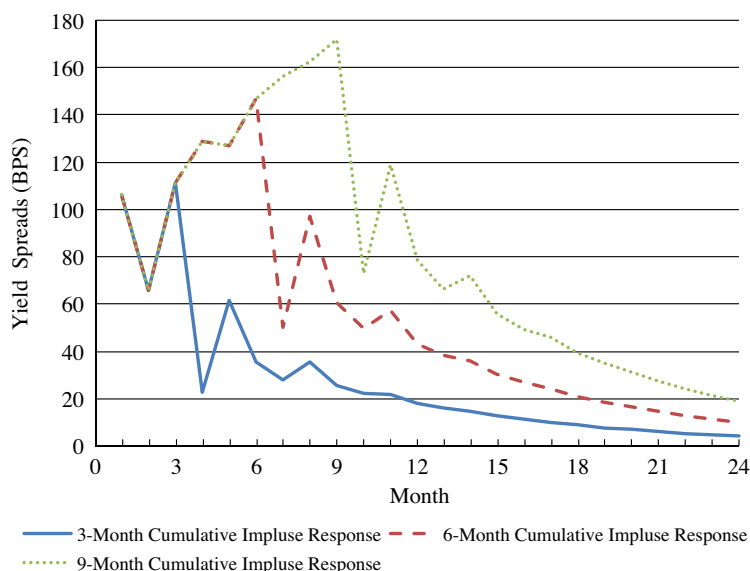


Fig. 3. Cumulative impulse response of Δ Subprime yield spread to \ln CDO issuance (1997.01–2006.08).

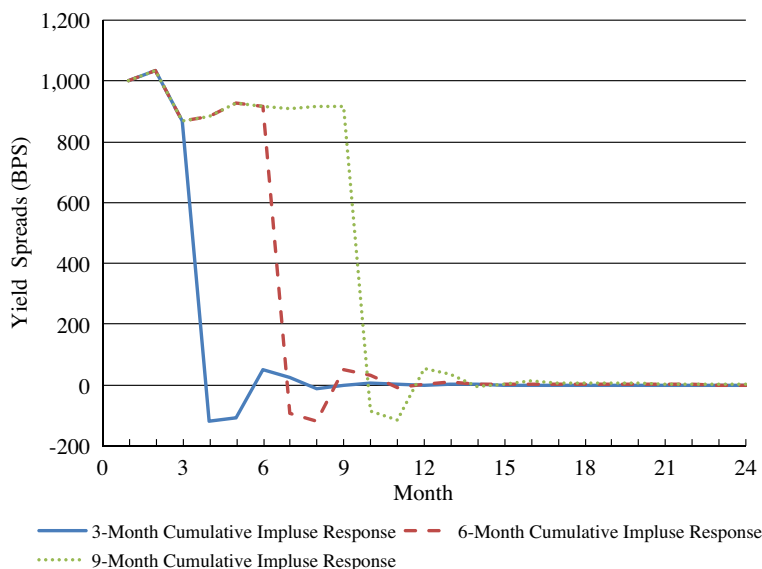


Fig. 4. Cumulative impulse response of Δ Subprime yield spread to Δ Subprime yield spread (1997.01–2006.08).

and more precisely estimated in the recent 2001–2006 period.

4.2. Vector autoregression and impulse responses

To further explore the dynamic dependency between subprime–Treasury yield spreads and the underlying time series variables, we adopt a structural vector autoregression (VAR) model to conduct the dynamic analysis. Our primary focus is to test how pricing of subprime mortgage-backed securities changes in response to shocks in the issuance of subprime-backed CDOs. We estimate the VAR model with changes in subprime–Treasury yield spread and natural log of issuance of subprime-backed CDOs as

endogenous variables, whereas the remaining control variables in Eq. (1) are treated as exogenous variables.

Column 1 of Table 6 reports the estimated impulse responses of the subprime–Treasury spread to the natural log of issuance of subprime-backed CDOs based on the full sample period of 1997.01–2006.08. Findings indicate significant and negative impulse response of the subprime MBS–Treasury spread to CDO issuance lagged by 2 months. Results also indicate a marginally significant and positive impulse response to CDO issuance lagged 1 month and a less significant and negative impulse response to the CDO issuance in current month, perhaps reflecting a mean-reverting effect in the impulse response estimates (correction due to over-reaction in the previous month’s impulse

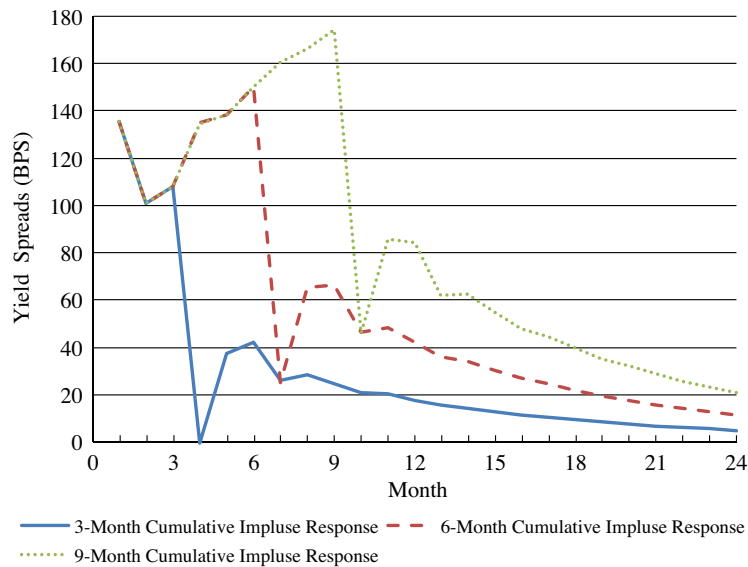


Fig. 5. Cumulative impulse response of Δ Subprime yield spread to \ln CDO issuance (2001.02–2006.08).

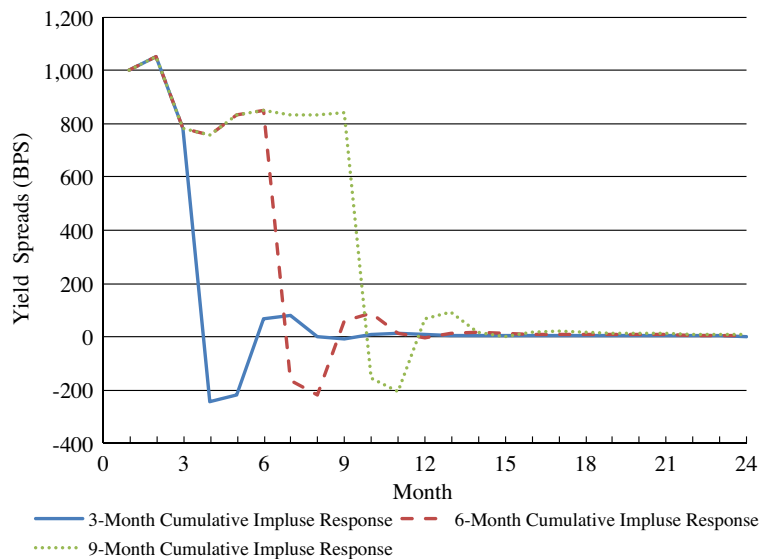


Fig. 6. Cumulative impulse response of Δ Subprime yield spread to Δ Subprime yield spread (2001.02–2006.08).

responses). Column 2 of Table 6 reports the estimated impulse response of CDO issuance to changes in subprime–Treasury spreads. As such, column 2 provides a test for potential reverse causality in the estimated relationship between CDO issuance and subprime–Treasury yield spreads. Results here indicate that none of the estimated lagged coefficients for changes in subprime MBS–Treasury spreads are statistically significant. Accordingly, findings from column 2 (consistent with results of Granger Tests reported above) indicate the absence of reverse causality in the relationship between subprime–Treasury spreads and CDO issuance. However, the significant and positive coefficients for CDO issuance lagged by 1 month and 2 months

suggest that momentum associated with positive supply shocks in CDO issuance might lead to further compression in the subprime MBS–Treasury spreads.¹⁷ While the estimated coefficients for the exogenous variables have the expected signs, most are smaller in magnitude relative to those reported in Table 5. Overall, VAR results provide important insight of dynamic process of CDO issuance effects in the

¹⁷ The last two columns in Table 6 report the estimated impulse responses associated with a sub-sample covering the 2001.02–2006.08 period. While the qualitative results remain similar to those reported above, the coefficient estimates are less significant, perhaps due in part to the limited set of data points in the shorter sampling period.

determination of subprime mortgage-backed securities spreads.¹⁸

4.3. Simulations

Finally, we use the results from Table 6 to simulate the impulse responses of changes in subprime mortgage-backed security pricing associated with 3-month, 6-month, and 9-month cumulative adverse shocks to the issuance of subprime-backed CDOs. Using the full sample (1997.01–2006.08) estimates, Fig. 3 reports the simulated impulse responses in the subprime MBS–Treasury yield spreads. Results indicate that a one standard deviation negative shock in ln CDO issuance lasting 3 months, 6 months, and 9 months, results in a widening of the subprime MBS–Treasury yield spread of 110 bps, 150 bps, and 170 bps, respectively. Further, findings reported in Fig. 4 suggest that a one standard deviation positive shock in subprime MBS–Treasury spreads lasting for 3–9 months leads to further widening in spreads by up to 1000 bps. Figs. 5 and 6 report the simulated impulse responses of subprime MBS–Treasury yield spreads based on estimates from the 2001.02–2006.08 sub-sample. Results are largely robust to the change in estimation period.

5. Conclusion

This research evaluates effects of the emergence of the CDO market on the pricing of subprime residential mortgage-backed securities. In so doing, it applies monthly-frequency time-series to evaluate the determinants of yield spreads of subprime residential mortgage-backed securities (MBS) to comparable-maturity Treasury bonds. Empirical tests suggest that factors associated with the termination risks of the underlying residential mortgage contracts, including proxies for mortgage prepayment and credit risk, importantly affect the magnitude of MBS–Treasury spreads. Spreads on mortgage-backed securities also appear sensitive to returns and related return volatility among alternative asset classes. Finally, holding constant the ratio of subprime-backed CDO to subprime MBS issuance, research indicates that the emergence and rapid capitalization of the CDO market was associated with a significant tightening of subprime RMBS–Treasury yield spreads. Results of VAR and other robustness tests serve to corroborate the results. Further, findings from dynamic simulations based on the VAR estimates indicate substantial widening of spreads in the wake of implosion in the CDO market. Research findings suggest the importance of innovations in derivative securities markets to the pricing of underlying subprime collateral. Results similarly indicate that the 2007 closure of the CDO market exerted substantial upward pressure on MBS spreads, and in so doing

contributed to changes in the pricing, underwriting and related demise of subprime mortgages.

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¹⁸ To test robustness of results, we also estimate the VAR with three endogenous variables, including subprime–Treasury yield spread, the natural log of issuance of subprime-backed CDOs, and the ratio of subprime-backed CDOs to subprime MBS issuance. Reported results (above) are largely robust to the inclusion of an additional endogenous variable.

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